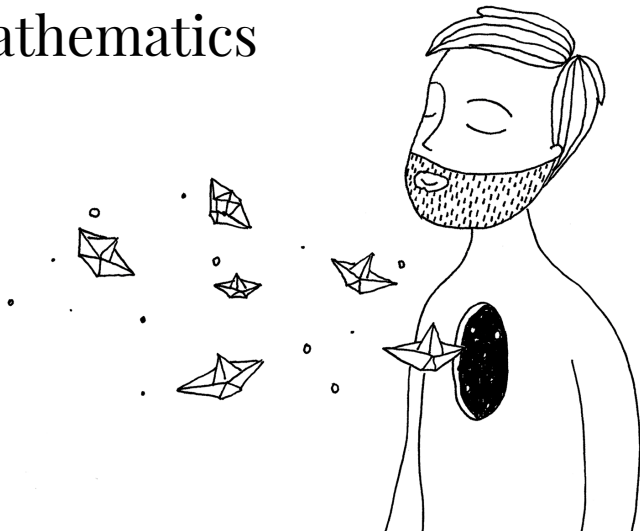


# 4509 – Bridging Mathematics

## Markov Chains

PAULO FAGANDINI



# Discrete-time Markov chains

1. Time is indexed by an integer variable, say  $n$ .
2. At period  $n$ , the **state** of the chain is denoted by  $X_n$ .
3.  $\mathcal{S}$  is a finite set of possible states, then  $X_n \in \mathcal{S}$ .
4. We will allow for  $m$  different states, then  $\mathcal{S} = \{1, 2, \dots, m\}$ , for  $m \in \mathbb{N}$ .

# Discrete-time Markov chains

## Definition

Markov Chain The Markov chain is described in terms of its **transition probabilities**  $p_{ij}$ : whenever the state happens to be  $i$ , there is probability  $p_{ij}$  that the next state is equal to  $j$ :

$$p_{ij} = P(X_{n+1} = j | X_n = i), \quad i, j \in \mathcal{S}$$

with  $p_{ij} \geq 0$  and  $\sum_{j=1}^m p_{ij} = 1 \quad \forall i$ .

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**Note:** the probability does not depend on time, nor anything else than the present state.

## How to specify then a Markov Model?

- Identify:

1.  $\mathcal{S}$  the set of states.
2. the set of possible transitions,  $(i, j)$  where  $p_{ij} > 0$
3. the values for those  $p_{ij}$

- The Markov chain specified by this model is a sequence of r.v.s  $X_0, X_1, X_2, \dots$ , that can take values in  $\mathcal{S}$ , and which satisfy:

$$P(X_{n+1} = j | X_n = i, \{X_\nu = i_\nu\}_{\nu=0}^{n-1}) = p_{ij}$$

for any  $n$ , and any  $i, j \in \mathcal{S}$ , and all possible sequences  $i_0, \dots, i_{n-1}$  of earlier states.

It is convenient to sort all these probabilities in a two-dimensional array like this:

$$\begin{bmatrix} p_{11} & p_{12} & \dots & p_{1m} \\ p_{21} & p_{22} & \dots & p_{2m} \\ \vdots & \vdots & \ddots & \vdots \\ p_{m1} & p_{m2} & \dots & p_{mm} \end{bmatrix}$$

This is called the **Transition Probability Matrix**. This matrix is defined as having in each row  $i$  and column  $j$  the probability of transitioning from state  $i$  to state  $j$ .

## Example, Bertsekas and Tsitsiklis (2008)

Alice is taking a probability class and in each week, she can be either up-to-date or she may have fallen behind. If she is up-to-date in a given week, the probability that she will be up-to-date (or behind) in the next week is 0.8 (or 0.2, respectively) . If she is behind in the given week, the probability that she will be up-to-date (or behind) in the next week is 0.6 (or 0.4, respectively) . We assume that these probabilities do not depend on whether she was up-to-date or behind in previous weeks, so the problem has the typical Markov chain character (the future depends on the past only through the present)

Let 1 be the state of being up-to-date and 2 that she fell behind.



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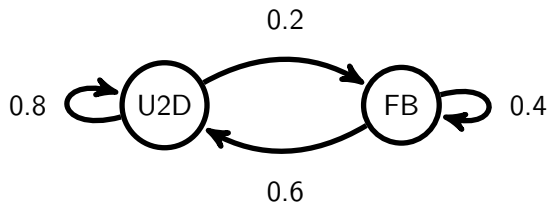
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The transition probability matrix:

$$\begin{bmatrix} 0.8 & 0.2 \\ 0.6 & 0.4 \end{bmatrix}$$

The transition probability graph:



We have said that the state today depends only on the state in the previous period. This is true, however, we can get around this constraint.

Consider the following example:

1. A working machine can be working the next day with probability  $p$ , and be broken with probability  $1 - p$ .
2. A broken machine can be working the next day with probability  $q$ , and remain broken with probability  $1 - q$ .

However, what happens if the machine cannot be fixed for, say, 4 straight days? Maybe we need to buy a new one. To model this we can introduce new states to our system.

1. A working machine can be working the next day with probability  $p$ , and be 1-day broken with probability  $1 - p$ , and zero for  $n$ -days broken for  $n > 1$ .
2. A 1-day broken machine can be working the next day with probability  $q$ , and become 2-day broken with probability  $1 - q$ , and zero for  $n$ -days broken for  $n \neq 2$ .
3. A 2-days broken machine can be working the next day with probability  $r$ , and become broken for 3 days with probability  $1 - r$ , and zero for  $n$ -days broken for  $n \neq 3$ .
4. A 3-days broken machine can be working the next day with probability  $s$ , and become broken for 4 days with probability  $1 - s$ , and zero for  $n$ -days broken for  $n \neq 4$ .
5. A 4-days broken machine can be working with probability 1, and zero for all the other broken states.

## Definition ( $n$ -Step Transition Probabilities)

Let  $r_{ij}(n)$  represent the probability that the state after  $n$  time periods will be  $j$ , given that the current state is  $i$ .

$$r_{ij}(n) = P(X_n = j | X_0 = i)$$

## Proposition (Chapman-Kolmogorov)

*The  $n$ -step transition probabilities can be generated by the recursive formula:*

$$r_{ij}(n) = \sum_{k=1}^m r_{ik}(n-1)p_{kj}, \quad \text{for } n > 1, \text{ and all } i, j$$

*starting with*

$$r_{ij}(1) = p_{ij}$$

Note that this is an element of the following matrix:

$$\begin{bmatrix} p_{11} & p_{12} & \dots & p_{1m} \\ p_{21} & p_{22} & \dots & p_{2m} \\ \vdots & \vdots & \ddots & \vdots \\ p_{m1} & p_{m2} & \dots & p_{mm} \end{bmatrix}^n$$

This “realization” allow us to be able to ask and answer some interesting questions:

- What can we say about limits? What happens as  $n \rightarrow \infty$ ?
- The dependence of the state at  $n$  over the initial state becomes smaller as  $n$  increases.
- What can we say qualitatively about the behavior of this markov chain?



Consider the transition matrix for the example we just saw:

$$A = \begin{bmatrix} 0.8 & 0.2 \\ 0.6 & 0.4 \end{bmatrix}$$

Note how as  $n \rightarrow \infty$   $r_{ij}(n)$  goes to a limit that does not depend on the initial state.

Consider the transition matrix for the example we just saw:

$$A = \begin{bmatrix} 0.8 & 0.2 \\ 0.6 & 0.4 \end{bmatrix} \quad A^2 = \begin{bmatrix} 0.7600 & 0.2400 \\ 0.7200 & 0.2800 \end{bmatrix}$$

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Note how as  $n \rightarrow \infty$   $r_{ij}(n)$  goes to a limit that does not depend on the initial state.

## Definition (Accessible state)

A state  $j$  is accessible from a state  $i$  if  $\exists n \in \mathbb{N}$  such that the  $n$ -step transition probability  $r_{ij}(n)$  is positive, i.e., if there is positive probability of reaching  $j$ , starting from  $i$ , after some number of periods.



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## Definition (Recurrent state)

Let  $A(i)$  be the set of states that are accessible from  $i$ . We say that  $i$  is **recurrent** if  $\forall j \in A(i) \Rightarrow i \in A(j)$ .

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## Definition (Transient state)

A state is called **transient** if it is not recurrent.

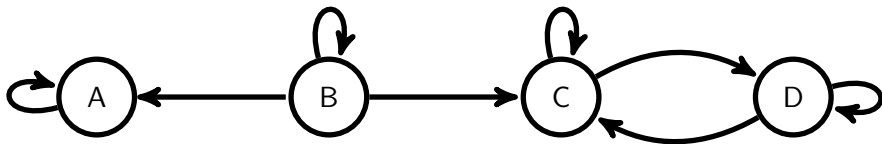
## Corollary

*A recurrent state will be visited an infinity amount of times.*

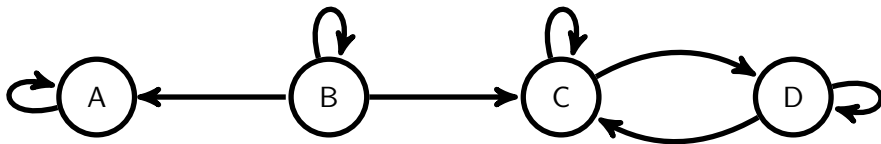
## Corollary

*A transient state will be visited a finite amount of times.*

Which of the following nodes are transient and which are recurrent?

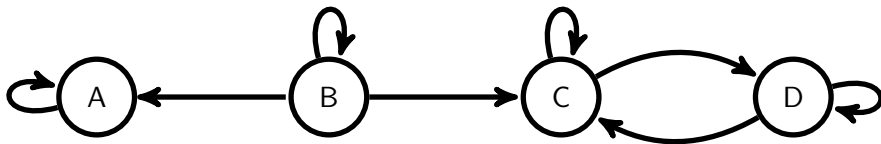


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Recurrent

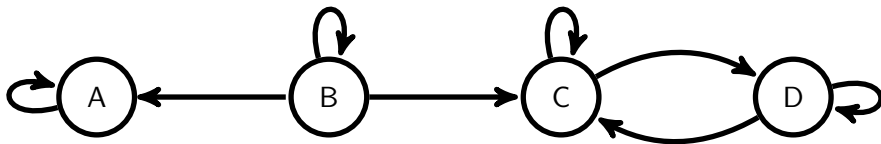
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Recurrent

Transient

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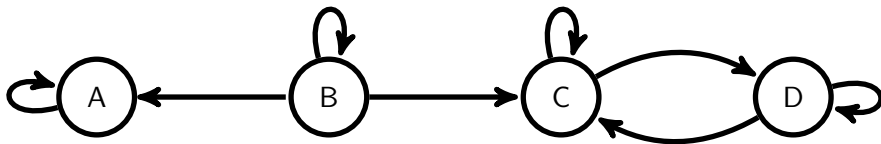


Recurrent

Transient

Recurrent

Which of the following nodes are transient and which are recurrent?



Recurrent

Transient

Recurrent

Recurrent



## Definition (Recurrent class)

If  $i$  is a recurrent state, the set of states  $A(i)$  that are accessible from  $i$  form a **recurrent class** (or simply a class), meaning that states in  $A(i)$  are all accessible from each other, and no state outside  $A(i)$  is accessible from them.

# Steady state behavior

When we talk about steady state in Markov Chains, it is not the “state” that is steady, but the probabilities of arriving to a certain state, remember the example we had before?

$$\pi_j = P(X_n = j), \quad \text{when } n \text{ is large.}$$

## Theorem (Steady-State Convergence Theorem)

*Consider a Markov chain with a single recurrent class, which is periodic. Then, the states  $j$  are associated with steady-state probabilities  $\pi_j$  that have the following properties:*

1. *For each  $j$ , we have*

$$\lim_{n \rightarrow \infty} r_{ij}(n) = \pi_j, \quad \forall i$$

2. *The  $\pi_j$  are the unique solution to the system of equations below:*

$$\pi_j = \sum_{k=1}^m \pi_k p_{kj}, \quad j = 1, \dots, m,$$

$$1 = \sum_{k=1}^m \pi_k$$

3. *We have*

$$\pi_j = 0 \quad , \text{ for all transient states } j$$

$$\pi_j > 0 \quad , \text{ for all recurrent states } j$$

Note that the steady-state probabilities add up to 1...

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Therefore these form a probability distribution on the state space, this is called the **stationary distribution** of the chain.

## Definition (Balance Equations)

The equations

$$\pi_j = \sum_{k=1}^m \pi_k p_{kj}, \quad j = 1, \dots, m,$$

are called **balance equations**, and they are a direct consequence of the first part of the Steady-State Convergence Theorem, and the Chapman-Kolmogorov equation.

## Definition (Normalization Equation)

The equation

$$\sum_{k=1}^m \pi_k = 1$$

is known as the **normalization equation**.

# Example

Consider our original example:  $p_{11} = 0.8$   $p_{12} = 0.2$   $p_{21} = 0.6$   $p_{22} = 0.4$   
Clearly, on the limit  $r_{ij} \rightarrow \pi_j$  if this converges, then the balance equations say:

$$\pi_1 = \pi_1 p_{11} + \pi_2 p_{21} \quad \pi_2 = \pi_1 p_{12} + \pi_2 p_{22}$$

Which, replacing, become:

$$\pi_1 = 0.8\pi_1 + 0.6\pi_2 \quad \pi_2 = 0.2\pi_1 + 0.4\pi_2$$

Solving, we obtain  $\pi_1 = 3\pi_2$  in both equations, which together with the normalization equation  $\pi_1 + \pi_2 = 1$  lead us to:

$$\pi_1 = 0.75$$

$$\pi_2 = 0.25$$